Swayam: Distributed Autoscaling to Meet SLAs of Machine Learning Inference Services with Resource Efficiency

Arpan Gujarati MPI-SWS, Germany arpanbg@mpi-sws.org Sameh Elnikety MSR, USA samehe@microsoft.com

Kathryn S. McKinley Google, Inc., USA mckinley@cs.utexas.edu

APPENDIX

Assumption 1. The average time to send a message from any frontend to any backend is d_1 and the average time to send a message from any backend to any frontend is d_2 . Requests arrive at an average rate of λ and that the average time to service a request is $1/\mu$.

DEFINITION 1. **RAND LB Policy:** Upon receipt of a request from the broker, the frontend immediately forwards it to a random backend. If the backend is busy, it sends back a busy response to the frontend. Upon receipt of a busy response, the frontend forwards the request to another random backend after Δ time units.

THEOREM 1 (WAITING TIME DISTRIBUTION). If there are n backends, and requests are assigned to backends as per the **RAND LB Policy**, then the expected p^{th} percentile waiting time of the requests can be approximated as:

$$\omega_p = d_1 + \left(\frac{\ln\left(1 - \frac{p}{100}\right)}{\ln\left(\frac{\lambda}{n\cdot\mu}\right)} - 1\right) \cdot (d_1 + d_2 + \Delta). \tag{1}$$

PROOF. Assume that $d_1, d_2, \Delta \ll 1/\mu$. As a result, the duration for which the system remains non-work-conserving is negligible. Thus, the probability that any backend is busy at any given point of time is given by the overall system utilization:

$$P_{busy} = \frac{\lambda}{n \cdot \mu}.$$
 (2)

Similarly, let the probability that any backend is idle at any given point of time be denoted as:

$$P_{idle} = 1 - P_{busy} = 1 - \frac{\lambda}{n \cdot \mu}.$$
(3)

Let P_r denote the probability that a request finds an idle backend after r retries, i.e., P_r denotes the probability that a request is assigned to busy backends during the first r attempts, and only in the $r + 1^{st}$ retry, the request is assigned to an idle backend. Thus,

$$P_r = (P_{busy})^r \cdot P_{idle}.$$
 (4)

Let W_r denote the waiting time of the request that finds an idle backend after r retries. W_r consists of r round-trip latencies for the first r attempts, r delays of time Δ each enforced by the frontend, and a single frontend-to-backend communication latency for the last successful attempt. Thus,

$$W_r = r \cdot (d_1 + d_2) + r \cdot \Delta + d_1$$

= $r \cdot (d_1 + d_2 + \Delta) + d_1.$ (5)

Yuxiong He MSR, USA yuxhe@microsoft.com

Björn B. Brandenburg MPI-SWS, Germany bbb@mpi-sws.org

Let ω_p denote the p^{th} percentile waiting time and suppose that it consists of delays due to r_{max} retries. Thus, $\omega_p = W_{r_{max}}$, i.e.,

$$\omega_p = r_{max} \cdot (d_1 + d_2 + \Delta) + d_1$$

$$\equiv r_{max} = \frac{\omega_p - d_1}{d_1 + d_2 + \Delta}.$$
 (6)

If any request requires less than or equal to r_{max} retries to find an idle backend, then its waiting time is also less than or equal to ω_p . Therefore:

$$\sum_{k=0}^{r_{max}} P_k = \frac{p}{100}.$$

= $\sum_{k=0}^{r_{max}} ((P_{busy})^k \cdot P_{idle}) = \frac{p}{100}.$
= $P_{idle} \cdot \sum_{k=0}^{r_{max}} (P_{busy})^k = \frac{p}{100}.$

{by using sum of geometric progression}

$$\equiv P_{idle} \cdot \frac{1 - (P_{busy})^{r_{max}+1}}{1 - P_{busy}} = \frac{p}{100}.$$

{from Eq. 3}

$$\equiv 1 - (P_{busy})^{r_{max}+1} = \frac{p}{100}.$$

{from Eq. 2}

$$= 1 - \left(\frac{\lambda}{n \cdot \mu}\right)^{r_{max}+1} = \frac{p}{100}.$$

$$= \left(\frac{\lambda}{n \cdot \mu}\right)^{r_{max}+1} = 1 - \frac{p}{100}.$$

$$= r_{max} + 1 = \frac{\ln\left(1 - \frac{p}{100}\right)}{\ln\left(\frac{\lambda}{n \cdot \mu}\right)}.$$

{from Eq. 6}

$$= \frac{\omega_p - d_1}{d_1 + d_2 + \Delta} + 1 = \frac{\ln\left(1 - \frac{p}{100}\right)}{\ln\left(\frac{\lambda}{n \cdot \mu}\right)}.$$

$$= \omega_p = d_1 + \left(\frac{\ln\left(1 - \frac{p}{100}\right)}{\ln\left(\frac{\lambda}{n \cdot \mu}\right)} - 1\right) \cdot (d_1 + d_2 + \Delta).$$

Hence, Eq. 1 is proved.

THEOREM 2 (RESPONSE TIME DISTRIBUTION). If there are n backends, if the requests are assigned to backends as per the **RAND LB Policy**, and the CDF of the request execution times is denoted by $CDF_{exe}(x)$, then the probability that the response time of a request is less than or equal to RT is given by

$$\sum_{r=0}^{r_{max}} \left(\frac{\lambda}{n \cdot \mu}\right)^r \cdot \left(1 - \frac{\lambda}{n \cdot \mu}\right) \cdot CDF_{exe}(RT - WT_r), \tag{7}$$

 $where \, r_{max} = \lfloor (RT - d_1) / (d_1 + d_2 + \Delta) \rfloor \, and \, WT_r = r \cdot (d_1 + d_2 + \Delta) + d_1.$

PROOF. Let *Y* and *Z* denote the random variables corresponding to request execution times and waiting times, respectively. Let R = Y + Z denote the random variable corresponding to the request response time. Thus, by the general formula for the distribution of the sum of two independent discrete variables:

$$P(R \le RT) = \sum_{z=-\infty}^{\infty} P(Z = z) \cdot P(Y \le RT - z)$$

{since *Z* can only take discrete values based on the number of retries, i.e., for *r* retries, $Z = WT_r = r \cdot (d_1 + d_2 + \Delta) + d_1$, and since *r* varies from 0 to ∞ }

$$P(R \le RT) = \sum_{r=0}^{\infty} P(Z = WT_r) \cdot P(Y \le RT - WT_r)$$

{since *R*, *Y*, *Z* are non-negative, $R \le RT$ implies that $Z = WT_r \le RT$, which in turn implies that $r \le r_{max} = \lfloor (RT - d_1)/(d_1 + d_2 + \Delta) \rfloor$

$$P(R \le RT) = \sum_{r=0}^{r_{max}} P(Z = WT_r) \cdot P(Y \le RT - WT_r)$$

{since $P(Z = WT_r)$ is equivalent to $P_r = (P_{busy})^r \cdot P_{idle}$ in Eq. 4, and since $P_{busy} = (\lambda)/(n \cdot \mu)$ and $P_{idle} = 1 - (\lambda)(n \cdot \mu)$ for the RAND LB Policy}

$$P(R \le RT) = \sum_{r=0}^{r_{max}} \left(\frac{\lambda}{n \cdot \mu}\right)^r \cdot \left(1 - \frac{\lambda}{n \cdot \mu}\right)$$
$$\cdot P(Y \le RT - WT_r)$$

{since $P(Y \le RT - WT_r)$ is equivalent to the CDF of the execution time distribution}

$$P(R \le RT) = \sum_{r=0}^{r_{max}} \left(\frac{\lambda}{n \cdot \mu}\right)^r \cdot \left(1 - \frac{\lambda}{n \cdot \mu}\right)$$
$$\cdot CDF_{exe}(RT - WT_r).$$

Hence, Eq. 7 is proved.